

Uncovering disassortativity in large scale-free networks

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Joint work with Remco van der Hofstad

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- ► signature log-log plot: $\log p_k = log(const) \alpha \log k$
- ► Faloutsos, Faloutsos, Faloutsos (1999): power laws in Internet

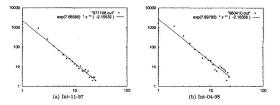


Figure 5: The outdegree plots: Log-log plot of frequency f_d versus the outdegree d.

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But Power Law is not everything!

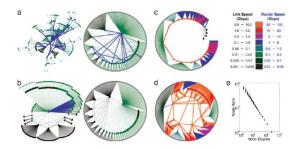
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- Albert, Jeong and Barabasi (2000): Achille's heel of Internet: Internet is sensitive to targeted attack
- Doyle et al. (2005): Robust yet fragile nature of Internet: Internet is not a random graph, it is designed to be robust



Example: Spread of infections

 Classical epidemiology, e.g. Adnerson and May (1991): epidemic only if infection rate exceeds a critical value

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Example: Technological versus economical networks



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- ▶ It is clearly important how the network is wired
- To start with: do hubs connect to each other? YES for banks, NO for Internet
- Assortative networks: nodes with similar degree connect to each other.
- Disassortative networks: nodes with large degrees tend to connect to nodes with small degrees.

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- Newman (2002): assortativity measure ρ_n

$$\rho_n = \frac{\frac{1}{|E|} \sum_{ij \in E} d_i d_j - \left(\frac{1}{|E|} \sum_{ij \in E} \frac{1}{2} (d_i + d_j)\right)^2}{\frac{1}{|E|} \sum_{ij \in E} \frac{1}{2} (d_i^2 + d_j^2) - \left(\frac{1}{|E|} \sum_{ij \in E} \frac{1}{2} (d_i + d_j)\right)^2}$$

 Statistical estimation of the correlation coefficient between degrees on two ends of a random edge

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- Statistical estimation of the correlation coefficient between degrees on two ends of a random edge
- Very popular measure of assortativity!

- Preferential Attachment graph appears to be assortatively neutral (Newman 2003, Dorogovtsev et al. 2010)
- Recent criticism: ρ_n depends on the size of the networks (Raschke et al. 2010; Dorogovtsev et al. 2010)

- ρ_n is a statistical estimation for the coefficient of variation $\rho = \frac{E(XY) [E(X)]^2}{Var(X)},$
- ► X and Y are the degrees of the nodes on the two ends of a randomly chosen edge

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$$Var(X)$$
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- ► Problems? YES!!!

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- ► Problems? YES!!!
- X and Y are power law r.v.'s, exponent $\alpha 1$

$$P(X = k) = kp_k/E(\text{degree}).$$

► In real networks (WWW) we often have 2 < α < 3, so $E(X) = \sum_{k} k \frac{kp_k}{E(\text{degree})} = \infty$

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- ► In real networks (WWW) we often have 2 < α < 3, so $E(X) = \sum_{k} k \frac{kp_k}{E(\text{degree})} = \infty$
- ρ is not defined in the power law model! Then: what are we measuring?

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Assortative and disassortative graphs

► Newman(2003)

	network	type	size n	assortativity r	error σ_r	ref.
$_{\rm social}$	physics coauthorship	undirected	52 909	0.363	0.002	a
	biology coauthorship	undirected	1520251	0.127	0.0004	a
	mathematics coauthorship	undirected	253 339	0.120	0.002	Ь
	film actor collaborations	undirected	449 913	0.208	0.0002	с
	company directors	undirected	7673	0.276	0.004	d
	student relationships	undirected	573	-0.029	0.037	e
	email address books	directed	16881	0.092	0.004	f
$technological \begin{cases} \\ \\ \\ \\ \\ \end{cases}$	power grid	undirected	4941	-0.003	0.013	g
	Internet	undirected	10697	-0.189	0.002	h
	World-Wide Web	directed	269504	-0.067	0.0002	i
	software dependencies	directed	3 162	-0.016	0.020	j
biological {	protein interactions	undirected	2115	-0.156	0.010	k
	metabolic network	undirected	765	-0.240	0.007	1
	neural network	directed	307	-0.226	0.016	m
	marine food web	directed	134	-0.263	0.037	n
	freshwater food web	directed	92	-0.326	0.031	0

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- Technological and biological networks are disassortative, ρ_n < 0
- Social networks are assortative, $\rho_n > 0$

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- Social networks are assortative, $\rho_n > 0$
- ► Note: large networks are never strongly disassortative...

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ρ_n in terms of moments of the degrees

► Write

$$\sum_{ij\in E} \frac{1}{2}(d_i + d_j) = \sum_{i\in V} d_i^2, \qquad \sum_{ij\in E} \frac{1}{2}(d_i^2 + d_j^2) = \sum_{i\in V} d_i^3$$

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$$\rho_n = \frac{\sum_{ij \in E} d_i d_j - \frac{1}{|E|} \left(\sum_{i \in V} d_i^2\right)^2}{\sum_{i \in V} d_i^3 - \frac{1}{|E|} \left(\sum_{i \in V} d_i^2\right)^2}.$$

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Theorem (Extreme value theory)

 $D_1, D_2, ..., D_n$ are i.i.d. with $1 - F(x) = P(D > x) = Cx^{-\alpha+1}$. Then

$$\lim_{n\to\infty} P\left(\frac{\max\{D_1, D_2, \dots, D_n\} - b_n}{a_n} \leqslant x\right) = \exp(-(1+\delta x)^{-1/\delta}),$$

with $\delta = 1/(\alpha - 1)$, $a_n = \delta C^{\delta} n^{\delta}$, $b_n = C^{\delta} n^{\delta}$. (Therefore, the maximum is 'of the order' $n^{1/(\alpha - 1)}$)

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Theorem (CLT for heavy tails)

 $D_1, D_2, ..., D_n$ are *i.i.d.* with $1 - F(x) = P(D > x) = Cx^{-\alpha+1}$. If $p > \alpha - 1$ then

$$\frac{1}{a_n}\sum_{i=1}^n X_i^p \stackrel{d}{\to} Z,$$

where $a_n = [1 - F]^{-1}(1/n^p) = C^{1/(\alpha-1)} n^{p/(\alpha-1)}$ and Z has a stable distribution with parameter $(\alpha - 1)/p$. (Therefore, the sum is 'of the order' $n^{p/(\alpha-1)}$)

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In the empirical setting

• $P(d_1 \ge x) \approx C x^{-\alpha+1}$

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- $P(d_1 \ge x) \approx C x^{-\alpha+1}$
- max{ d_1, d_2, \ldots, d_n } = $O(n^{1/(\alpha-1)})$
- ► Alternative interpretation for the maximum: $P(d \ge x) = 1/n \Rightarrow x = O(n^{1/(\alpha-1)})$

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► $\mathbb{P}(d_i = k) = p_k = const \cdot k^{-\alpha}$, usually $\alpha \in (2, 4)$
► If $p > \alpha - 1$ then $\mathbb{E}(D^p) = \infty$
► CLT: for $p > \alpha - 1$ holds
 $\frac{1}{n} \sum_{i \in V} d_i^p \sim c_p n^{p/(\alpha-1)-1}$,

► But we get the same result just by adding up k^pp_k from k = 1 to k = n^{1/(α-1)}.

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$$cn \leq |E| \leq Cn, (SLLN)$$

$$cn^{1/(\alpha-1)} \leq \max_{i \in [n]} d_i \leq Cn^{1/(\alpha-1)},$$

$$cn^{\max\{p/(\alpha-1),1\}} \leq \sum_{i \in [n]} d_i^p \leq Cn^{\max\{p/(\alpha-1),1\}}, \quad p = 2, 3,$$

where C, c > 0.

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where C, c > 0. Very natural and non-restrictive assumptions for power law graphs.

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Back to ρ_n

$$\rho_n = \frac{\text{crossproducts} - \text{expectation}^2}{\text{variance}} \geqslant - \frac{\text{expectation}^2}{\text{variance}} = \rho_n^-$$

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• We have
$$\sum_{i \in V} d_i^3 \ge c n^{3/(\alpha-1)}$$

But also

$$\frac{1}{|E|} \left(\sum_{i \in V} d_i^2\right)^2 \leq (C^2/c) n^{\max\{4/(\alpha-1)-1,1\}}$$

When α ∈ (2, 4) we have max{4/(α − 1) − 1, 1} < 3/(α − 1), so that the denominator of ρ_n[−] outweighs its numerator.

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► Take e.g.
$$\alpha = 2.5$$

► $4/(\alpha - 1) - 3/(\alpha - 1) = -1/3$
► $\rho_n^- = O(n^{-1/3})$

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- ▶ ρ_n^- converges to zero as $n \to \infty$ in ANY power law graph
- Large scale-free graphs are never disassortative!
- Reason: high variability in values \Rightarrow dependence on *n*

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Alternative: rank correlations

• $((X_i, Y_i))_{i=1}^n$ random variables

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- $((X_i, Y_i))_{i=1}^n$ random variables
- r_i^X and r_i^Y the rank of X_i and Y_i , respectively

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Alternative: rank correlations

- $((X_i, Y_i))_{i=1}^n$ random variables
- r_i^X and r_i^Y the rank of X_i and Y_i , respectively
- ► Spearman's rho:

$$\rho_n^{\text{rank}} = \frac{\sum_{i=1}^n (r_i^X - (n+1)/2)(r_i^Y - (n+1)/2)}{\sqrt{\sum_{i=1}^n (r_i^X - (n+1)/2)^2 \sum_i^n (r_i^Y - (n+1)/2)^2}}$$

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- ▶ r_i^X and r_i^Y are from uniform distribution: $n \cdot Uniform(0, 1)$

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- Correlation coefficient for r_i^X and r_i^Y
- ▶ r_i^X and r_i^Y are from uniform distribution: $n \cdot Uniform(0, 1)$
- ► Factor *n* cancels, no influence of high dispersion

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H. Hotelling and M.R. Pabst (1936):

'Certainly where there is complete absence of knowledge of the form of the bivariate distribution, and especially if it is believed not to be normal, the rank correlation coefficient is to be strongly recommended as a means of testing the existence of relationship.'

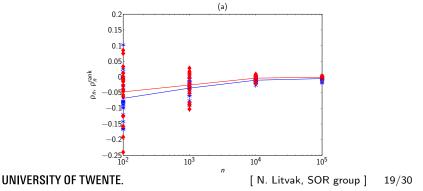
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Configuration model (CM)

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- ► d_i 's are original degrees in the CM, $\ell_n = \sum_i d_i$. In CMIE we obtain:

$$\rho_n = \frac{2\sum_{i \in V} 2d_i - \frac{1}{2\ell_n} \left(\sum_{i \in V} d_i^2 + 2\ell_n\right)^2}{\sum_{i \in V} d_i^3 + 4\ell_n - \frac{1}{2\ell_n} \left(\sum_{i \in V} d_i^2 + 2\ell_n\right)^2}.$$

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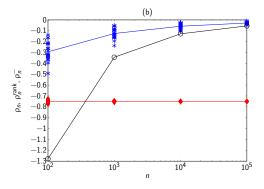
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Configuration model with intermediate edge: results

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N. Litvak, SOR group] 21/30

Preferential Attachment (PA) graph

- Albert and Barabási (1999), simplest version with one outgoing edge per node.
- Nodes arrive one at a time. A new node connects to a node i with probability proportional to current degree of i.

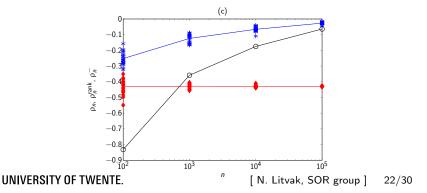
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$$\rho_n = \frac{\sum_{ij \in E} d_i d_j - \frac{1}{|E|} \left(\sum_{i \in V} d_i^2\right)^2}{\sum_{i \in V} d_i^3 - \frac{1}{|E|} \left(\sum_{i \in V} d_i^2\right)^2}.$$

Two possible scenarios:

- Denominator outweighs numerator, $\rho_n \rightarrow 0$
- Denominator and numerator are of the same order of magnitude. Limit?

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Collection of bipartite graphs

►
$$((X_i, Y_i))_{i=1}^n$$
 i.i.d.

 $X = bU_1 + bU_2$, $Y = bU_1 + aU_2$, b > 0, a > 1

 U_1 , U_2 i.i.d. random variables with power law tail, exponent α .

- For i = 1, ..., n, we create a complete bipartite graph of X_i and Y_i vertices, respectively.
- ► These *n* complete bipartite graphs are not connected to one another.
- Extreme scenario of a network consisting of highly connected clusters of different size. Such networks can serve as models for physical human contacts and are used in epidemic modelling (Eubank et al. 2004).
- ► Disassortative for n = 1 but positive dependence between X and Y prevails for larger n.

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Collection of bipartite graphs: analysis

►
$$|V| = \sum_{i=1}^{n} (X_i + Y_i), |E| = 2 \sum_{i=1}^{n} X_i Y_i,$$

$$\sum_{i \in V} d_i^p = \sum_{i=1}^{n} (X_i^p Y_i + Y_i^p X_i) \qquad \sum_{ij \in E} d_i d_j = 2 \sum_{i=1}^{n} (X_i Y_i)^2.$$

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► Take $\mathbb{P}(U_j > x) = c_0 x^{-\alpha+1}$, where $c_0 > 0$, $x \ge x_0$, and $\alpha \in (4, 5)$, so that $E[U^3] < \infty$, but $E[U^4] = \infty$.

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► Then
$$|E|/n \xrightarrow{p} 2E[XY] < \infty$$
 and
 $\frac{1}{n} \sum_{i \in V} d_i^2 \xrightarrow{p} E[XY(X+Y)] < \infty.$

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Collection of bipartite graphs: analysis

Theorem (L& van der Hofstad, 2012)

$$n^{-4/(\alpha-1)}b^{-4}\sum_{i=1}^{n}(X_{i}^{3}Y_{i}+Y_{i}^{3}X_{i}) \xrightarrow{d} (a^{3}+a)Z_{1}+2Z_{2},$$

 $n^{-4/(\alpha-1)}b^{-4}\sum_{i=1}^{N}(X_{i}Y_{i})^{2} \xrightarrow{d} a^{2}Z_{1}+Z_{2},$

where Z_1 and Z_2 and two independent stable distributions with parameter $(\alpha - 1)/4$.

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[N. Litvak, SOR group] 26/30

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Result:

$$ho_n \stackrel{d}{
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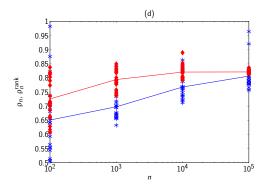
which is a random variable taking values in $(2a/(1+a^2), 1)$, a > 1.

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[N. Litvak, SOR group] 26/30

Collection of bipartite graphs: results

 ρ_n (blue), ρ_n^{rank} (red), and mean ρ_n^- (black) in 20 simulations for different n



[N. Litvak, SOR group] 27/30

Dataset	Description	# nodes	max d	ρ _n	ρ_n^{rank}	ρ_n^-
stanford-cs	web domain	9,914	340	-0.1656	-0.1627	-0.4648
eu-2005	.eu web crawl	862,664	68,963	-0.0562	-0.2525	-0.0670
uk@100,000	.uk web crawl	100,000	55,252	-0.6536	-0.5676	-1.117
uk@1,000,000	.uk web crawl	1,000,000	403,441	-0.0831	-0.5620	-0.0854
enron	e-mailing	69,244	1,634	-0.1599	-0.6827	-0.1932
dblp-2010	co-authorship	326,186	238	0.3018	0.2604	-0.7736
dblp-2011	co-authorship	986,324	979	0.0842	0.1351	-0.2963
hollywood-2009	co-starring	1,139,905	11,468	0.3446	0.4689	-0.6737

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[N. Litvak, SOR group] 28/30

[N. Litvak, SOR group] 29/30

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[N. Litvak, SOR group] 29/30

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 - In a graph the degrees on the ends of random edges are in general dependent. Can we analyse Spearman's rho? Work in progress.

Thank you!

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[N. Litvak, SOR group] 30/30